## 23 Random real zeroes: no derivatives

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## **23a** Random element of $L_2[0,1]$

Continuing Sect. 22d, we consider a Gaussian process

$$\Xi: [0,1] \to G \subset L_2(\Omega, P), \quad \Xi(t) = f_1(t)g_1 + f_2(t)g_2 + \dots,$$

where  $(g_1, g_2, ...)$  is an orthonormal basis of G, and  $f_k(t) = \langle \Xi(t), g_k \rangle$  are measurable. Necessarily,<sup>1</sup>

$$\forall t \in [0,1] \quad |f_1(t)|^2 + |f_2(t)|^2 + \dots = ||\Xi(t)||^2 < \infty.$$

We upgrade  $\Xi$  to the corresponding random element of  $L_0[0, 1]$  (as explained in Sect. 22d), denoted by  $X : \Omega \to L_0[0, 1]$ . In general,  $\int_0^1 ||\Xi(t)||^2 dt = \sum_k \int |f_k(t)|^2 dt$  need not be finite. From now on we assume that it is:

$$\int_0^1 \|\Xi(t)\|^2 \,\mathrm{d}t < \infty \,;$$

then, by Tonelli's theorem,

$$\mathbb{E} \int_0^1 |X(t)|^2 \, \mathrm{d}t = \int_0^1 \left( \mathbb{E} \, |X(t)|^2 \right) \, \mathrm{d}t = \int_0^1 \|\Xi(t)\|^2 \, \mathrm{d}t < \infty \,,$$

which shows that X is in fact a random element of  $L_2[0, 1]$ . We approximate X by another random element  $X_n$  of  $L_2[0, 1]$ ,

$$X_n(t) = g_1 f_1(t) + \dots + g_n f_n(t) \,.$$

We may also treat X and  $X_n$  as elements of  $L_2([0,1] \times \Omega)$ .

**23a1 Exercise.**  $X_n \to X$  in  $L_2([0,1] \times \Omega)$ .<sup>2</sup>

Prove it.

<sup>&</sup>lt;sup>1</sup>This is also sufficient (think, why).

<sup>&</sup>lt;sup>2</sup>In fact, almost surely the series converges in  $L_2(0, 1)$ .

**23a2 Exercise.** For every  $f \in L_2[0,1]$  the random variables  $\langle f, X_n \rangle = \langle f_1, f \rangle g_1 + \dots + \langle f_n, f \rangle g_n$  converge (as  $n \to \infty$ ) in  $L_2(\Omega)$  to the random variable  $\langle f, X \rangle = \int_0^1 f(t)X(t) dt$ .

Prove it.

Thus,

$$\operatorname{Var}\langle f, X \rangle = \sum_{k} |\langle f, f_k \rangle|^2 \le C ||f||^2$$

for some  $C \le \sum_k ||f_k||^2 = \int_0^1 ||\Xi(t)||^2 dt < \infty$ .

**23a3 Proposition.** Let C be such that

$$\forall f \in L_2[0,1] \quad \operatorname{Var}\langle f, X \rangle \le C \|f\|^2 \,.$$

Let  $\psi : L_2[0,1] \to \mathbb{R}$  be a Lip(1) function. Then the random variable  $\psi(X)$  belongs to GaussLip( $\sqrt{C}$ ).

First, we need the duality argument used already in 11c3.

**23a4 Lemma.**  $||a_1f_1+a_2f_2+\ldots||^2 \leq C(a_1^2+a_2^2+\ldots)$  for all  $(a_1,a_2,\ldots) \in l_2$ .

Proof. We introduce a linear operator  $S : l_2 \to L_2[0,1]$  by  $Sa = \sum a_k f_k$ ; the series converges in  $L_2[0,1]$ , since  $\sum ||a_k f_k|| = \sum |a_k| \cdot ||f_k|| \le (\sum |a_k|^2)^{1/2} (\sum ||f_k||^2)^{1/2} < \infty$ . We have  $\forall a \in l_2 \quad \forall f \in L_2[0,1] \quad \langle f, Sa \rangle = \langle S^*f, a \rangle$ , where  $S^* : L_2[0,1] \to l_2$ ,  $S^*f = (\langle f, f_1 \rangle, \langle f, f_2 \rangle, \ldots)$ . We note that  $\operatorname{Var}\langle f, X \rangle = ||S^*f||^2$ ; thus,  $||S^*f||^2 \le C||f||^2$  for all f.

We note that  $\operatorname{Var}\langle f, X \rangle = \|S^*f\|^2$ ; thus,  $\|S^*f\|^2 \leq C\|f\|^2$  for all f. Finally,

$$||Sa|| = \sup_{\|f\| \le 1} \langle f, Sa \rangle = \sup_{\|f\| \le 1} \langle S^*f, a \rangle \le \sup_{\|f\| \le 1} ||S^*f|| ||a|| \le \sqrt{C} ||a||.$$

Proof of the proposition. Similarly to the proof of 22d5 we assume that  $(\Omega, P) = (\mathbb{R}^{\infty}, \gamma^{\infty}), g_k$  are the coordinates, and will prove that  $\psi(X)$  is a  $\operatorname{Lip}(\sqrt{C})$  function on  $(\mathbb{R}^{\infty}, \gamma^{\infty})$ .

We take  $n_1 < n_2 < \ldots$  such that  $\sum_{i=1}^{n_k} f_i g_i \to X$  (as  $k \to \infty$ ) almost everywhere on  $[0, 1] \times \Omega$ .

 $\square$ 

<sup>&</sup>lt;sup>1</sup>In fact,  $n_k = k$  fit.

Given  $a \in l_2$ , we introduce  $h = a_1 f_1 + a_2 f_2 + \cdots \in L_2[0, 1]; \|h\|^2 \leq C \|a\|^2$ by 23a4. For almost all  $(t, x) \in [0, 1] \times (\mathbb{R}^{\infty}, \gamma^{\infty})$  we have

$$X(x+a,t) - X(x,t) = \lim_{k} \sum_{i=1}^{n_{k}} (x_{i}+a_{i})f_{i}(t) - \lim_{k} \sum_{i=1}^{n_{k}} x_{i}f_{i}(t) =$$
$$= \lim_{k} \sum_{i=1}^{n_{k}} a_{i}f_{i}(t) = h(t).$$

Thus, X(x+a) - X(x) = h for almost all  $x \in (\mathbb{R}^{\infty}, \gamma^{\infty})$ . Finally,

$$|\psi(X(x+a)) - \psi(X(x))| \le ||X(x+a) - X(x)|| = ||h|| \le \sqrt{C} ||a||.$$

Here is a useful formula for the variance:

(23a5) 
$$\operatorname{Var}\langle f, X \rangle = \int_0^1 \int_0^1 f(s) f(t) \left( \mathbb{E} \,\Xi(s) \Xi(t) \right) \,\mathrm{d}s \,\mathrm{d}t$$

for every  $f \in L_2[0, 1]$ . Proof:

$$\begin{split} \mathbb{E}\left(\int f(t)X(t)\,\mathrm{d}t\right)^2 &= \mathbb{E}\,\iint f(s)X(s)f(t)X(t)\,\mathrm{d}s\mathrm{d}t = \\ &= \iint \left(\mathbb{E}\,f(s)X(s)f(t)X(t)\right)\mathrm{d}s\mathrm{d}t\,, \end{split}$$

since

$$\mathbb{E} \iint |f(s)X(s)f(t)X(t)| \,\mathrm{d}s \mathrm{d}t = \mathbb{E} \left( \int |f(t)X(t)| \,\mathrm{d}t \right)^2 \leq \\ \leq \mathbb{E} \left( \int |f(t)|^2 \,\mathrm{d}t \right) \left( \int |X(t)|^2 \,\mathrm{d}t \right) = \|f\|^2 \int_0^1 \|\Xi(t)\|^2 \,\mathrm{d}t < \infty \,.$$

# **23b** Using assumption $A_n$

Let  $\Xi : \mathbb{R} \to G \subset L_2(\Omega, P)$  be a mean-square continuous stationary Gaussian random process on  $\mathbb{R}$ , and  $\mu$  its spectral measure:

$$\mathbb{E} \Xi(0)\Xi(t) = \int_{-\infty}^{+\infty} e^{i\lambda t} \mu(d\lambda) = \int_{-\infty}^{+\infty} \cos \lambda t \, \mu(d\lambda) \, .$$

<sup>&</sup>lt;sup>1</sup>In fact, the distribution  $X[\gamma^{\infty}]$  of X is a Gaussian measure on  $L_2[0,1]$ , and h is its admissible shift.

Here is another useful formula for the variance, this time in terms of the spectral measure (recall 11c4):

(23b1) 
$$\operatorname{Var}\langle f, X \rangle = \int \left| \int_0^1 f(t) \mathrm{e}^{\mathrm{i}\lambda t} \, \mathrm{d}t \right|^2 \mu(\mathrm{d}\lambda)$$

for every  $f \in L_2[0, 1]$ . Proof:

$$\begin{aligned} \operatorname{Var}\langle f, X \rangle &= \iint f(s) f(t) \Big( \int e^{i\lambda(t-s)} \mu(\mathrm{d}\lambda) \Big) \, \mathrm{d}s \mathrm{d}t = \\ &= \int \mu(\mathrm{d}\lambda) \Big( \int f(s) \overline{e^{i\lambda s}} \, \mathrm{d}s \Big) \Big( \int f(t) e^{i\lambda t} \, \mathrm{d}t \Big) \,, \end{aligned}$$

since

$$\int \mu(\mathrm{d}\lambda) \iint \left| f(s)f(t)\mathrm{e}^{\mathrm{i}\lambda(t-s)} \right| \mathrm{d}s\mathrm{d}t = \mu(\mathbb{R}) \left( \int \left| f(t) \right| \mathrm{d}t \right)^2 < \infty.$$

We generalize assumptions A and  $A_n$  of Sect. 2 as follows. ASSUMPTION A:

$$\mu(\mathbb{R}) = 1.$$

That is,  $X(0) \sim N(0, 1)$ . Otherwise we may rescale X. ASSUMPTION  $A_n$ : assumption A holds, and in addition,<sup>1</sup>

$$\forall \lambda \in [0,\infty) \quad \mu([\lambda,\lambda+1]) \leq \frac{1}{n}.$$

The argument of Sect. 11c still applies, recall (11c5): for every  $f \in L_2[0,1]$ ,

$$\int |g|^2 \,\mathrm{d}\mu \le C \Big( \int |g(\lambda)|^2 \,\mathrm{d}\lambda \Big) \sup_{\lambda} \mu \big( [\lambda, \lambda + 1] \big) ;$$

as before,  $g(\lambda) = \int_0^1 e^{i\lambda t} f(t) dt$ ,  $||g||_2^2 = 2\pi ||f||_2^2$ , and

$$\operatorname{Var}\langle f, X \rangle = \int |g|^2 \,\mathrm{d}\mu \,.$$

Thus, assumption  $A_n$  implies (recall 11c3)

$$\operatorname{Var}\langle f, X \rangle \le \frac{C}{n} \|f\|^2,$$

<sup>&</sup>lt;sup>1</sup>Alternatively you may take  $\lambda \in \mathbb{R}$ ; it is the same up to a factor 2 absorbed by an absolute constant.

and, by 23a3,

$$\psi(X) \in \operatorname{GaussLip}(C/\sqrt{n})$$

whenever  $\psi: L_2[0,1] \to \mathbb{R}$  is a Lip(1) function.

Now all arguments of 11d, 11e apply, and so, Theorems 2a2, 2a3 are generalized as follows.

Let X be a jointly measurable modification of a mean-square continuous stationary Gaussian random process on  $\mathbb{R}$ , satisfying assumption  $A_n$ .

**23b2 Proposition.** Let a function  $\varphi : \mathbb{R} \to \mathbb{R}$  be continuous almost everywhere, and

$$\sup_{x} \frac{|\varphi(x)|}{1+|x|} < \infty \,.$$

Then the random variable

$$\xi = \int_0^1 \varphi \big( X(t) \big) \, \mathrm{d}t$$

is integrable,  $\mathbb{E}\xi = \int \varphi \, \mathrm{d}\gamma^1$ , and for every  $\varepsilon > 0$ ,

$$\mathbb{P}\left(\left|\xi - \mathbb{E}\,\xi\right| \ge \varepsilon\right) \le 2\mathrm{e}^{-c_{\varepsilon,\varphi}n}$$

for some  $c_{\varepsilon,\varphi} > 0$  (dependent on  $\varepsilon$  and  $\varphi$  only, not on n).

#### 23b3 Proposition.

$$\mathbb{P}(T(X(\cdot)) \ge \varepsilon) \le 2\mathrm{e}^{-c_{\varepsilon}n}$$

for some  $c_{\varepsilon} > 0$  dependent on  $\varepsilon$  only.

As before, for  $f \in L_1[0, 1]$ ,

$$T(f) = \inf_{g} \int_{0}^{1} |f(t) - g(t)| \, \mathrm{d}t$$

where the infimum is taken over all measurable  $g : (0,1) \to \mathbb{R}$  that send Lebesgue measure to  $\gamma^1$ .

A trivial rescaling of t by arbitrary L > 0 turns assumption  $A_n$  and Proposition 23b2 into the following.

ASSUMPTION  $A_{n,L}$ : assumption A holds, and in addition,

$$\forall \lambda \in [0,\infty) \quad \mu\left(\left[\lambda,\lambda+\frac{1}{L}\right]\right) \leq \frac{1}{n}.$$

$$\xi = \frac{1}{L} \int_0^L \varphi \left( X(t) \right) dt$$

is integrable,  $\mathbb{E}\xi = \int \varphi \, d\gamma^1$ , and for every  $\varepsilon > 0$ ,

$$\mathbb{P}(|\xi - \mathbb{E}\xi| \ge \varepsilon) \le 2\mathrm{e}^{-c_{\varepsilon,\varphi}n}$$

for some  $c_{\varepsilon,\varphi} > 0$ .

Now, at last, we can deal with a single process, getting rid of assumption  $A_{n,L}$ .

**23b5 Theorem.** Let X be a jointly measurable<sup>1</sup> modification of a meansquare continuous stationary Gaussian random process on  $\mathbb{R}$  whose spectral measure has a bounded density.<sup>2</sup> Let a function  $\varphi : \mathbb{R} \to \mathbb{R}$  be continuous almost everywhere, and

$$\sup_{x} \frac{|\varphi(x)|}{1+|x|} < \infty \,.$$

Then random variables

$$\xi_L = \frac{1}{L} \int_0^L \varphi(X(t)) \, \mathrm{d}t \quad \text{for } L \in (0,\infty)$$

are integrable,  $\mathbb{E}\xi_L = \int \varphi \, \mathrm{d}\gamma^1$ , and for every  $\varepsilon > 0$ ,

$$\mathbb{P}\left(\left|\xi_L - \mathbb{E}\,\xi_L\right| \ge \varepsilon\right) \le 2\mathrm{e}^{-c_{\varepsilon,\varphi,M}L}$$

for some  $c_{\varepsilon,\varphi,M} > 0$  (dependent only on  $\varepsilon$ ,  $\varphi$  and the supremum M of the spectral density, not on L).

23b6 Exercise. Prove Theorem 23b5.

23b7 Exercise. Formulate and prove a single-process counterpart of 23b3.

## 23c Dimension two, and higher

A two-component (in other words,  $\mathbb{R}^2$ -valued) Gaussian random process on a set T may be defined as a pair  $(\Xi_1, \Xi_2)$  of Gaussian processes  $\Xi_1, \Xi_2 : T \to G \subset L_2(\Omega, P)$ . Or equivalently, as a Gaussian process  $\Xi : T \times \{1, 2\} \to$ 

<sup>&</sup>lt;sup>1</sup>Sample continuity is of course sufficient (by 22d3).

<sup>&</sup>lt;sup>2</sup>Equivalently,  $\sup_{a < b} \frac{\mu([a,b])}{b-a} < \infty$ .

 $G.^1$  Similarly, a two-component random function  $\xi$  on T is a pair  $(\xi_1, \xi_2)$  of random functions  $\xi_1, \xi_2 : \Omega \to \mathbb{R}^T$ , or a random function  $\xi : \Omega \to \mathbb{R}^{T \times \{1,2\}} = \mathbb{R}^T \times \mathbb{R}^T$ . Clearly,  $(\xi_1, \xi_2)$  is a modification of  $(\Xi_1, \Xi_2)$  if and only if both  $\xi_1$  is a modification of  $\Xi_1$  and  $\xi_2$  is a modification of  $\Xi_2$ . Continuity and measurability properties are defined evidently.

The covariance function of  $\Xi : T \times \{1, 2\} \to G$  is  $(s, k; t, l) \mapsto \mathbb{E} \Xi(s, k) \Xi(t, l) = \mathbb{E} \Xi_k(s) \Xi_l(t)$ . Stationarity (assuming  $T = \mathbb{R}$ ) is, by definition (recall 21e1),

$$\forall s, t \in \mathbb{R} \ \forall k, l \in \{1, 2\} \ \mathbb{E} \Xi_k(s) \Xi_l(t) = \mathbb{E} \Xi_k(0) \Xi_l(t-s) \,.$$

For a stationarity  $\Xi : \mathbb{R} \times \{1, 2\} \to G$  the covariance function  $R : \mathbb{R} \times \{1, 2\} \times \{1, 2\} \to \mathbb{R}$  is, by definition,

$$R(t,k,l) = R_{k,l}(t) = \mathbb{E} \Xi_k(0) \Xi_l(t);$$

it determines the process up to isometry. Another function  $r : \mathbb{R} \to \mathbb{R}$ ,

$$r(t) = \mathbb{E} \langle \Xi(0), \Xi(t) \rangle = \mathbb{E} \left( \Xi_1(0) \Xi_1(t) + \Xi_2(0) \Xi_2(t) \right) = R_{1,1}(t) + R_{2,2}(t) ,$$

containing only a partial information about R, will be called the traced covariance function. Normalizing the process to r(0) = 1 one may call r the correlation function. However, such normalization is sometimes inconvenient, since the case  $\Xi(0) \sim \gamma^2$  leads to r(0) = 2.

Clearly, the function r is positive definite. Assuming mean square continuity of  $\Xi$  we apply Bochner's theorem and get the traced spectral measure,<sup>2</sup> — a symmetric measure  $\mu$  on  $\mathbb{R}$  such that

$$\mathbb{E} \langle \Xi(0), \Xi(t) \rangle = r(t) = \int e^{i\lambda t} \mu(d\lambda) \,.$$

In the finite-dimensional case treated in 11f,  $r(t) = \sum_k |a_k|^2 \cos \lambda_k t$  ( $a_k$  being vectors), thus,  $\mu = \sum_k |a_k|^2 (\delta_{\lambda_k} + \delta_{-\lambda_k})/2$ .

Similarly to 23a we upgrade a two-component process  $\Xi$  to the corresponding random element<sup>3</sup> X of  $L_2([0,1] \to \mathbb{R}^2)$  and consider

$$\langle f, X \rangle = \langle f_1, X_1 \rangle + \langle f_2, X_2 \rangle$$

<sup>&</sup>lt;sup>1</sup>A coordinate-free definition of a *E*-valued Gaussian process on *T*, for a finitedimensional linear space *E*, may be given as follows: it is a linear map from  $E^*$  to  $G^T$ .

<sup>&</sup>lt;sup>2</sup>The full (non-traced) spectral measure may be treated as a matrix-valued measure on  $\mathbb{R}$ , or equivalently, a 2 × 2 matrix whose elements are (signed) measures on  $\mathbb{R}$ . For an *E*-valued process one gets a "scalar product" on  $E^*$  whose values are (signed) measures on  $\mathbb{R}$ .

<sup>&</sup>lt;sup>3</sup>Just upgrade  $\Xi_1$  to  $X_1$ ,  $\Xi_2$  to  $X_2$ , and take  $X = (X_1, X_2)$ .

for  $f = (f_1, f_2) \in L_2([0, 1] \to \mathbb{R}^2)$ . We cannot calculate  $\operatorname{Var}\langle f, X \rangle$  in terms of the *traced* spectral measure  $\mu$  (like (23b1)), but we can bound it:<sup>1</sup>

$$\begin{aligned} \operatorname{Var}\langle f, X \rangle &\leq 2 \int \left| \int_0^1 f(t) \mathrm{e}^{\mathrm{i}\lambda t} \, \mathrm{d}t \right|^2 \mu(\mathrm{d}\lambda) = \\ &= 2 \int \left( \left| \int_0^1 f_1(t) \mathrm{e}^{\mathrm{i}\lambda t} \, \mathrm{d}t \right|^2 + \left| \int_0^1 f_2(t) \mathrm{e}^{\mathrm{i}\lambda t} \, \mathrm{d}t \right|^2 \right) \mu(\mathrm{d}\lambda) \,. \end{aligned}$$

Proof:

$$\begin{aligned} \operatorname{Var}\langle f, X \rangle &= \|\langle f, X \rangle\|^2 = \|\langle f_1, X_1 \rangle + \langle f_2, X_2 \rangle\|^2 \le 2\|\langle f_1, X_1 \rangle\|^2 + 2\|\langle f_2, X_2 \rangle\|^2 \\ &= 2\int \left| \int_0^1 f_1(t) \mathrm{e}^{\mathrm{i}\lambda t} \, \mathrm{d}t \right|^2 \mu_{1,1}(\mathrm{d}\lambda) + 2\int \left| \int_0^1 f_2(t) \mathrm{e}^{\mathrm{i}\lambda t} \, \mathrm{d}t \right|^2 \mu_{2,2}(\mathrm{d}\lambda) \,, \end{aligned}$$

where  $\mu_{1,1}$  is the spectral measure for  $X_1$ , and  $\mu_{2,2}$  — for  $X_2$ ; it remains to note that  $\mu = \mu_{1,1} + \mu_{2,2}$  (think, why).

Assumption A is replaced with

$$\Xi(0) \sim \gamma^2$$

(which implies  $\mu(\mathbb{R}) = 2$ ); assumption  $A_n$  still adds

$$\forall \lambda \in [0, \infty) \quad \mu([\lambda, \lambda + 1]) \le \frac{1}{n}$$

where  $\mu$  is the traced spectral measure. As before we get

$$\forall f \in L_2([0,1] \to \mathbb{R}^2) \quad \operatorname{Var}\langle f, X \rangle \leq \frac{C}{n} \|f\|^2;$$
$$\psi(X) \in \operatorname{GaussLip}(C/\sqrt{n})$$

whenever  $\psi : L_2([0,1] \to \mathbb{R}^2) \to \mathbb{R}$  is a Lip(1) function. Similarly to 11f, Propositions 23b2 and 23b3 generalize to two-component processes satisfying assumption  $A_n$ . Also Theorem 23b5 generalizes to two-component processes whose *traced* spectral measures have bounded densities.

All said about  $\mathbb{R}^2$  holds equally well for  $\mathbb{R}^d$ ,  $d = 3, 4, \ldots$ 

### 23d Hints to exercises

23b6: L = Cn.

 $<sup>^{1}</sup>$ In fact, the coefficient "2" is superfluous (see 11f for the discrete case); however, the stronger inequality is harder to prove.