Growth Effects of the Exchange-Rate Regime and the Capital-Account Openness in A Crises-Prone World Market: A Nuanced View

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June 2004

What is this section about:

This paper studies the direct and indirect effects of Macro policies pointing out the role of the indirect channels in evaluating the local effects.

Motivation:

Exchange rate regimes and capital account openness, influence output growth through two channels:

- 1. Directly, through their effect on the trade and financial sectors.
- 2. Indirectly, through their impact on the probability of a balance-of-payments crisis.

Usually, in offsetting directions. For instance, switching from float to peg:

- 1. Direct effect: positive effect on growth.
- 2. Indirect effect: negative by increasing the probability of a real exchange rate crisis.

Implications:

- 1. The empirical analysis demonstrates that the balance-of-payments policies significantly affect the probability of crises; and that the crisis probability, in turn, negatively affects output growth. Namely, the indirect effects of balance-of-payments policy on growth are substantial.
- 2. By controlling for the crisis probability in the growth equation, the paper uncovers also the direct channel. The direct effects of balance-of-payments policies are also demonstrated to be substantial. Indeed, in benchmark OLS regressions, which abstract from the effect on growth of crisis probabilities, effects of the balance-of-payments policy on output growth turn out to be negligible.
- 3. The indirect channel is intrisincally non linear: there exists a range for the exogenous variables for which the effect of policy on the likelihood of a crisis is large; and another range, where the effect is rather small.
- 4. As a consequence of the non linearity of the crisis probability

- function, overall growth effects of balance-of-payments policies depend crucially on the level of the crisis probability.
- **5.** Thus, the cost-benefit evaluation of balance-of-payments policies depends on the special characteristics of the economy.
- **6.** We propose a re-examination of discrete high inflation crises. Traditionally the analysis focuses on periods when inflation is above some threshold. Growth falls sharply during discrete high inflation crises.
- 7. By introducing a probability of internal (high inflation) crises, along with the probability of external (balance-of-payment) crisis that depends on the exchange rate system, we can further discern the non linear affect of the exchange rate system on growth.

A simple statistical model

Let $Y_{1,j,t}$ denote the growth rate of country j in time t as measured in terms of GDP per capita (growth rates).

Let $Y_{2,j,t}^*$ denote a latent variable indicating a threshold state of the economy: If $Y_{2,j,t}^* > 0$ a currency crises occurs; if $Y_{2,j,t}^* \leq 0$ a currency crises does not occurs. That is:

(1)

$$Y_{2,j,t} = \begin{cases} 1 & if Y_{2,j,t}^* \ge 0 \\ 0 & else \end{cases}$$

whereas $Y_{2,j,t}$ is a binary variable which equals 1 if currency crises occurs in country j at time t.

We assume that two policies: (i) a float-peg policy and (ii) a liberalization-controls policy. To simplify assume that policy decision is binary. Denote by D_1 the float-peg policy and by D_2 the liberalization-controls policy:

(2.a)

$$D_{1,j,t} = \begin{cases} 1 & \text{if peg} \\ 0 & \text{if float} \end{cases}$$

and:

(2.b)

$$D_{2,j,t} = \begin{cases} 1 & \text{if capital controls} \\ 0 & \text{if liberalization} \end{cases}$$

Therefore we could write the system as follows:

(3.a)

$$Y_{1,j,t} = \beta_1 X_{j,t} + \gamma_1 D_{1,j,t} + \delta_1 D_{1,j,t} + \phi_1 Y_{2,j,t} + \theta_1 Z_{1,j,t} + \varepsilon_{1,j,t}$$

and

(3.b)

$$Y_{2,j,t}^* = \beta_2 X_{j,t} + \gamma_2 D_{1,j,t} + \delta_2 D_{1,j,t} + \phi_2 Y_{1,j,t} + \theta_2 Z_{2,j,t} + \varepsilon_{2,j,t}$$

Whereas $X_{j,t}$ is a vector of country specific exogenous variables and $\varepsilon_{i,j,t}$ is a country specific i.i.d. random shock.

Let $P_{j,t}$ be the conditional probability that country j will face currency crisis in period t. Given our assumption,

(4)

$$P_{j,t} = \Pr(Y_{2,j,t} = 1 \mid \bullet) = \Phi(\beta_2 X_{j,t} + \gamma_2 D_{1,j,t} + \delta_2 D_{1,j,t} + \phi_2 Y_{1,j,t} + \theta_2 Z_{2,j,t} + \varepsilon_{2,j,t})$$

where Φ is the cdf of the unit normal distribution (above).

Note that we can identify the parameters of the "crisis-selection" equation by estimating the following Probit equation:

(5)

$$Y_{2,j,t} = \Phi(\beta_2 X_{j,t} + \gamma_2 D_{1,j,t} + \delta_2 D_{1,j,t} + \phi_2 Y_{1,j,t} + \theta_2 Z_{2,j,t} + \varepsilon_{2,j,t})$$

Where the projected likelihood for:

(6)

$$\hat{Y}_{2,j,t} = \Phi(\hat{\beta}_2 X_{j,t} + \hat{\gamma}_2 D_{1,j,t} + \hat{\delta}_2 D_{1,j,t} + \hat{\phi}_2 Y_{1,j,t} + \hat{\theta}_2 Z_{2,j,t})$$

Incorporating the probability of balance-of-payments crises

The estimated growth effect of D_1 and D_2 when the likelihood of a currency crisis is ignored are:

(7.a)

$$E(\hat{\gamma}_{1}^{IV}) = \frac{\partial E(Y_{1,j,t} \mid X_{j,t}, D_{1,j,t}^{IV}, D_{1,j,t}^{IV}, Z_{1,j,t})}{\partial D_{1,j,t}} = \gamma_{1} + \phi_{1} \frac{\partial \hat{Y}_{2,j,t}}{\partial D_{1,j,t}}$$

and:

(7.b)

$$E(\hat{\delta}_1^{IV}) = \frac{\partial E(Y_{1,j,t} \mid X_{j,t}, D_{1,j,t}^{IV}, D_{1,j,t}^{IV}, Z_{1,j,t})}{\partial D_{2,j,t}} = \delta_1 + \phi_1 \frac{\partial \hat{Y}_{2,j,t}}{\partial D_{2,j,t}}$$

where $D_{1,i,t}^{IV}$, $D_{1,i,t}^{IV}$ are the instrumented policy actions.

It is common wisdom that the likelihood of currency crisis has a negative effect on growth:

(8.a)

$$\phi_1 < 0$$

It is also common to assume that a peg exchange rate increases the likelihood of a currency crisis (all other things equal), and that capital controls reduce the probability of such a crisis:

(8.b)

$$\frac{\partial \hat{Y}_{2,j,t}}{\partial D_{1,i,t}} > 0$$

$$\frac{\partial \hat{Y}_{2,j,t}}{\partial D_{1,j,t}} > 0$$

$$\frac{\partial \hat{Y}_{2,j,t}}{\partial D_{1,j,t}} < 0$$

Implications:

(9.a)

$$E(\hat{\gamma}_1^{IV}) = \gamma_1 + \phi_1 \frac{\partial \hat{Y}_{2,j,t}}{\partial D_{1,j,t}} < \gamma_1 > 0$$

and:

(9.b)

$$E(\hat{\delta}_1^{IV}) = \delta_1 + \phi_1 \frac{\partial \hat{Y}_{2,j,t}}{\partial D_{2,j,t}} > \delta_1 < 0$$

Table 1: The Frequency of Crises, Switches Between Float and Peg and Switches between Capital Controls and Liberalizations (%)

| Variable | Frequency |
|-----------------------------|-----------|
| Crsises | 22.61 |
| Switches to peg | 1.71 |
| Switches to float | 3.91 |
| Switches to controls | 1.03 |
| Switches to liberalizations | 0.9 |

Table 2: List of Countries

| (4) | A I | (54) | Malausi |
|------|------------------------|-------|------------------|
| (1) | Algeria | (51) | Malawi |
| (2) | Argentina | (52) | Malaysia |
| (3) | Bangladesh | (53) | Maldives |
| (4) | Barbados | (54) | Mali |
| (5) | Belize | (55) | Malta |
| (6) | Benin | (56) | Mauritania |
| (7) | Bhutan | (57) | Mauritius |
| (8) | Bolivia | (58) | Mexico |
| (9) | Botswana | (59) | Morocco |
| (10) | Brazil | (60) | Myanmar |
| (11) | Burkina Faso | (61) | Nepal |
| (12) | Burundi | (62) | Nicaragua |
| (13) | Cameroon | (63) | Niger |
| (14) | Cape Verde | (64) | Nigeria |
| (15) | Central African | (65) | Oman |
| (16) | Chad | (66) | Pakistan |
| (17) | Chile | (67) | Panama |
| (18) | China | (68) | Papua New Guinea |
| (19) | Colombia | (69) | Paraguay |
| (20) | Comoros | (70) | Peru |
| (21) | Congo | (71) | Philippines |
| (22) | Cote d'Ivoire | (72) | Portugal |
| (23) | Dominican Rep. | (73) | Romania |
| (24) | Ecuador | (74) | Rwanda |
| (25) | Egypt, Arab Rep | (75) | Sao Tome and Pr |
| (26) | El Salvador | (76) | Senegal |
| (27) | Equatorial Guin | (77) | Seychelles |
| (28) | Ethiopia | (78) | Sierra Leone |
| (29) | Fiji | (79) | Solomon Islands |
| (30) | Gabon | (80) | Somalia |
| (31) | Gambia, The | (81) | South Africa |
| (32) | Ghana | (82) | Sri Lanka |
| (33) | Grenada | (83) | St. Vincent |
| (34) | Guatemala | (84) | Sudan |
| (35) | Guinea | (85) | Swaziland |
| (36) | Guinea-Bissau | (86) | Syrian Arab Rep |
| (37) | Guyana | (87) | Tanzania |
| (38) | Haiti | (88) | Thailand |
| (39) | Honduras | (89) | Togo |
| (40) | Hungary | (90) | Trinidad and To |
| (41) | India | (91) | Tunisia |
| (42) | Indonesia | (92) | Turkey |
| (43) | Iran, Islamic R | (93) | Uganda |
| (44) | Jamaica | (94) | Uruguay |
| (45) | Jordan | (95) | Vanuatu |
| (46) | Kenya | (96) | Venezuela |
| (47) | Lao PDR | (97) | Western Samoa |
| (48) | Lesotho | (98) | Zaire |
| (49) | Liberia | (99) | Zambia |
| (50) | Madagascar | (100) | Zimbabwe |
| · -/ | 5 - 2 - 2 - 2 | ` / | - |

Table 3: Exchange Regime and Capital Controls: Cyclical Effects

| Panel A: Dependent Variable: Growth Rates | OLS | OLS | FE | FE |
|---|-----------------------|-----------------------|-----------------------|-----------------------|
| Variables | (i) | (ii) | iii) | (iv) |
| Switching to peg between t-2 to t-1 | 1.6423 (0.7503)* | 4.6209 (1.4795)** | 1.2041 (0.9958) | 5.0215 (1.7630)** |
| Switching to float between t-2 to t-1 | 0.1761 (0.6483) | 0.6383 (0.6692) | -0.0539 (0.7039) | 0.2005 (0.7401) |
| Switching to Capital Controls between t-2 to t-1 | -1.8832 (0.8616)* | -4.7173 (1.5363)** | -1.9592 (1.0495) | -6.3843 (2.0713)** |
| The probability of having currency crisis this year ^ | | -9.6164 (5.0663) | | -12.7791 (4.9934)* |
| Controllers | | | | |
| 1970 GDP per capita | -0.0012 (0.0005)* | -0.0011 (0.0005)* | | |
| Currency crisis at time t-1 | 0.5612 (0.5949) | 2.7602 (1.2740)* | 0.7579 (0.4506) | 2.5482 (0.8331)** |
| Currency crisis at time t-2 | -2.1345 (0.6375)** | -1.5347 (0.7221)* | -1.6442 (0.4525)** | -2.2155 (0.4852)** |
| Growth rate at time t-1 | 0.2540 (0.0464)** | 0.2552 (0.0469)** | 0.1802 (0.0275)** | 0.2267 (0.0312)** |
| Growth rate at time t-2 | 0.1093 (0.0366)** | 0.1048 (0.0372)** | 0.0069 (0.0274) | -0.0224 (0.0313) |
| Panel B: Dependent Variable: Currency Crisis (0, | 1). 1 if REE(t) | -REE(t-1)>15 | % - Probit (dF | -/dX) estimators |
| 1970's GDP per capita | | 0.0000 (0.0000) | | |
| Switching to peg between t-2 to t-1 | | 0.3125 (0.0991)** | | 0.2893 (0.1028)** |
| Switching to float t-2 to t-1 | | 0.0557 (0.0510) | | 0.0325 (0.0516) |
| Switching to Capital Controls between t-2 to t-1 | | -0.2656 (0.0470)** | | -0.3313 (0.0524)** |
| Currency crisis at time t-1 | | 0.2299 (0.0377)** | | 0.1314 (0.0349)** |
| Currency crisis at time t-2 | | 0.0563 (0.0296) | | -0.0307 (0.0256) |
| Government def t-1 ^^ | | 0.0000 (0.0000) | | 0.0000 (0.0000) |
| Country fixed-effects | | No | | Yes |

Note:

Data includes 106 countries in the years 1970 to 1997

[^] Currency crisis =1 if the real exchange rate increased by 15% between t-1 to t (1 STD)

All specifications include linear time trend

⁽⁾ Standard errors in parenthesis

^{*} significant at 5%; ** significant at 1%

Table 4: Exchange Regime and Capital Controls: Cyclical and Persistent Effects

| Panel A: Dependent Variable: Growth Rates | | |
|---|------------------------|---------------------------|
| Variables | OLS (i) | FE (ii) |
| Peg at time t-1 | -0.6088 (0.2899)* | -0.1813 (0.4787) |
| Switching to peg between t-2 to t-1 | 3.9786 (1.2935)** | 4.9046 (1.4604)** |
| Switching to float between t-2 to t-1 | 0.4657 (0.7124) | 0.8090 (0.8382) |
| Capital Controls at t-1 | -1.2843 (0.4539)** | -1.1997 (0.9385) |
| Switching to Capital Controls between t-2 to t-1 | -1.2843 (0.4539)** | -5.9101 (1.7511)** |
| The probability of having currency crisis this year ^ | -7.9131 (6.0140) | -13.7764 (4.4409)** |
| Controllers | | |
| 1970 GDP per capita | -0.0013 (0.0006)* | |
| Currency crisis at time t-1 | 2.3069 (1.4183) | 2.6221 (0.7543)** |
| Currency crisis at time t-2 | -1.7389 (0.7269)* | -2.3438 (0.4911)** |
| Growth rate at time t-1 | 0.2481 (0.0456)** | 0.2247 (0.0312)** |
| Panel B: Dependent Variable: Currency Crisis (0,1). 1 i | if REE(t)-REE(t-1)>15% | - Probit (dF/dX) estimato |
| 1970's GDP per capita | 0.0000 (0.0000) | |
| Peg at time t-1 | -0.0192 (0.0221) | 0.0368 (0.0361) |
| Switching to peg between t-2 to t-1 | 0.2798 (0.1029)** | 0.2106 (0.1070)* |
| Switching to float t-2 to t-1 | 0.0801 (0.0567) | 0.1085 (0.0674) |
| Capital Controls at t-1 | -0.0383 (0.0283) | -0.1021 (0.0639) |
| Switching to Capital Controls between t-2 to t-1 | -0.2491 (0.0513)** | -0.2820 (0.0646)** |
| Currency crisis at time t-1 | 0.2264 (0.0373)** | 0.1255 (0.0345)** |
| Country fixed-effects | No | Yes |

Note:

Data includes 106 countries in the years 1970 to 1997

All specifications include linear time trend

[^] Currency crisis =1 if the real exchange rate increased by 15% between t-1 to t (1 STD)

⁽⁾ Standard errors in parenthesis

^{*} significant at 5%; ** significant at 1%

Table 5.a: The Frequency of Sudden Stop and Domestic Prices Crises Using Reinhart-Rogoff (2004) Classification*,**

| | | | Domestic Price Crises | | | |
|------------------------|---|---|-----------------------|------|-------|--|
| | | | 0 | 1 | | |
| Sudden Stops Crises | 0 | I | 24.6 | 9.9 | 34.5 | |
| CHSes | 1 | I | 29.3 | 36.3 | 65.5 | |
| | | | 53.9 | 46.1 | 100.0 | |

Notes:

We aggregate it into 2 main categories: (i) peg_rr, including the first 3 and (ii) float_rr, including the other two.

Domestic prices crisis = 1 if the inflation rate is above 20% per year and 0 otherwise.

Sudden stop crisis = 1 if the real exchange rate depreciation is above 15% per year and 0 otherwise.

^{*} Reinhart and Rogoff (2002) classified into 5 categories: (i) peg,

⁽ii) limited flexibility, (iii) managed floating, (iv) freely floating and (v) freely falling.

^{**} Data includes 58 countries in the years 1970 to 1997

Table 5.b: Switches Between Float and Peg Using Reinhart-Rogoff (2004) Classification*,**

| Variable | Frequency |
|-------------------|-----------|
| Switches to peg | 10.18 |
| Switches to float | 9.97 |

Notes:

^{*} Reinhart and Rogoff (2002) classified into 5 categories: (i) peg, (ii) limited flexibility, (iii) managed floating, (iv) freely floating and (v) freely falling. We aggregate it into 2 main categories: (i) peg_rr, including the first 3 and (ii) float_rr, including the other two.

^{**} Data includes 58 countries in the years 1970 to 1997

Table 6: Exchange Regime and Capital Controls Using Reinhart-Rogoff (2004) Classification*,** Fixed-Effects Estimators

Dependent Variable: Growth Rates

| Variables | (i) | (ii) | (iii) |
|--|--------------------|-------------------|--|
| Peg at time t-1 | 1.656 | 1.330 | 1.729 |
| | (0.557) | (0.549) | (0.565) |
| Capital Controls at t-1 | -0.439 | -0.587 | 0.156 |
| | (0.890) | (0.991) | (1.022) |
| Switching to Capital Controls between t-2 to t-1 | -5.852 | -3.374 | -6.155 |
| | (1.799) | (1.518) | (1.809) |
| The probability of having currency crisis this year^ excluding the effect of price crisis The probability of having currency crisis this year - real^^ including the effect of price crisis | -14.843 (4.937) | -6.824 (4.084) | -22.359 (7.996) 7.632 (6.578) |
| <u>Controllers</u> | | , | , , |
| Growth rate at time t-1 | 0.176 | 0.191 | 0.183 |
| | (0.034) | (0.034) | (0.034) |
| Growth rate at time t-2 | 0.008 | 0.022 | 0.019 |
| | (0.035) | (0.035) | (0.035) |
| Currency crisis at time t-1 | 2.812 | 0.917 | 3.340 |
| | (0.978) | (0.629) | (1.069) |
| Currency crisis at time t-2 | -1.904 | -1.804 | -1.831 |
| | (0.479) | (0.483) | (0.481) |
| Price (CPI) crisis at time t-1 | -0.100 | 1.078 | -1.251 |
| | (0.491) | (0.772) | (1.133) |
| Price (CPI) crisis at time t-2 | 0.385 | 0.374 | 0.468 |
| | (0.488) | (0.491) | (0.490) |

Notes:

^{*} Reinhart and Rogoff (2002) classified into 5 categories: (i) peg, (ii) limited flexibility, (iii) managed floating

^{, (}iv) freely floating and (v) freely falling. We aggregate it into 2 main categories: (i) peg_rr, including the first 3 and (ii) float_rr, including the other two.

^{**} Data includes 58 countries in the years 1970 to 1997

[^] The estimated the likelihood for a currency crisis ignoring the effect of price crisis.

[^] The estimated probability for a currency crisis including the effect of past price crisis

All specifications include linear time trend

⁽⁾ Standard errors in parenthesis