Estimating the effect of tax reform in differentiated product oligopolistic markets

Chaim Fershtman\textsuperscript{a, b}, Neil Gandal\textsuperscript{a, c, *}, Sarit Markovich\textsuperscript{d}

\textsuperscript{a}Tel Aviv University, 69978 Tel Aviv, Israel
\textsuperscript{b}CentER, Tilburg University, Tilburg, The Netherlands
\textsuperscript{c}CEPR, London, UK
\textsuperscript{d}University of Chicago, Chicago, IL, USA

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Abstract

The incidence of taxation has been extensively discussed in the public finance literature but mainly within a competitive market setting or within a homogenous good (Cournot type) oligopoly. In a differentiated product oligopoly, the effect of taxation can be more complex as the rate of taxation may affect not only the prices, but also the profile and quality of products that are sold in the market. In this paper, we examine the effects of changing tax regimes in a differentiated product oligopoly. In order to illustrate our approach, we employ data from one such market: the automobile market in Israel. © 1999 Elsevier Science S.A. All rights reserved.

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JEL classification: H2; L8

1. Introduction

The incidence of taxation and the design of an optimal tax system have been extensively discussed in the Public Finance literature; most of the work on this subject has been conducted within the framework of a competitive market or

*Corresponding author. Tel.: +972-3-640-6742; fax: +972-3-640-7382.
E-mail address: gandal@econ.tau.ac.il (N. Gandal)

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within the framework of a homogenous good oligopoly characterized by Cournot competition. Auerbach (1985) provides a good theoretical survey on commodity taxation in competitive markets. Katz and Rosen (1985), Myles (1987), and Stern (1987) examine commodity taxation issues such as tax design, optimal taxes and tax incidence in oligopolies in which firms compete on quantities and sell homogeneous products. In this paper we examine, both theoretically and empirically, the issues of tax reform in a very relevant market: oligopoly markets in which firms sell differentiated products.

In an oligopoly, in which products are differentiated both horizontally and vertically, the effect of taxation may be complex since the tax regime may affect not only prices, but also the profile and quality of the products that each firm sells. In this paper, we examine these issues in an oligopolistic market with horizontally and vertically differentiated products in which firms (may) sell multiple brands. In particular, the paper investigates the effect of changes in the taxation regime on equilibrium prices, tax incidence, the size and composition of the market, consumer surplus, government tax revenues, and firm profits. While we apply our framework using the automobile market in Israel, the type of analysis we employ is applicable to other differentiated product oligopolies.

In order to illustrate the complex effect of taxation on a differentiated good oligopoly, consider the automobile industry. Let us assume that in the relevant market, there are several firms that sell automobiles. Each of the firms sells different models. The models are vertically differentiated in that they differ in size, quality, features etc. At any given point in time these firms compete in prices. In determining the price for a specific model, each firm must take into account competition from other firms as well as the effect of the price of a particular model on the sales of its other models. Now assume that a tax is imposed. This will change the market equilibrium and will affect not just the prices of cars but also the equilibrium distribution of cars that each firm sells. For example, it is possible, that a uniform increase in the tax rate, while reducing overall sales in the market and overall sales of a particular firm, may increase sales of particular model. Such an effect will occur if, for example, as a result of the tax increase there is a shift in demand to smaller cars.

In our analysis, we employ recent advances in estimating discrete-choice models of product differentiation. These techniques, developed by Berry (1994) and Berry et al. (1995), enable structural estimation of both the demand and oligopoly pricing aspects that characterize differentiated products. The model of oligopolistic competition that we employ is based on Berry (1994) and is similar to the one used by Verboven (1996) and the one used in our previous work (Fershtman and Gandal, 1998).\(^1\) We use the model and data on prices, quantities, and characteris-

\(^1\)In Fershtman and Gandal (1998), we examined the effect of the Arab economic boycott on the Israeli automobile industry and estimated the peace dividend (in this industry) associated with the elimination of the boycott.
tics of the Israeli automobile market in 1994 and 1995 (the latest years for which the data are available) to estimate the demand and cost parameters for this market.

Given the estimated parameters, we analyze the effects of different tax regimes by simulating the oligopolistic market equilibrium under such regimes. The importance of the simulation is that it allows firms to adjust prices, as well as allowing for changes in quantity demanded.\(^2\)

An ad valorem tax regime is currently in place in Israel. In the paper, we examine the effect of changes in the percentage tax rate, as well as the effect of two alternative tax regimes: (i) a regime that consists of a per unit tax; and (ii) a tax based on engine size, as is common in several European countries (see Verboven, 1996). By comparing these regimes with the uniform percentage tax regime, we illustrate the product differentiation and market power effects that arise from changes in tax regimes.

We show that a change in the tax regime from an ad valorem tax to a per unit tax regime leads to significant reductions in the sales of small automobiles and significant increases in the sales of large automobiles. The change in composition is due to vertical product differentiation among the models. This demand effect is somewhat mitigated by market power. Firms take advantage of the increase in demand for larger cars (from the change in tax regime) by lowering prices by less than the reduction in taxes.

We are aware of two purely theoretical papers that examine the effect of taxation in differentiated product oligopolies. Gruenspecht (1988) considers the effect of export subsidies in an oligopoly in which firms sell horizontally differentiated products and compete on prices. In his model, price reductions do not bring new customers into the market; they only lead to market diversion from other competitors. Our setting explicitly allows for both effects and we measure them.

Cremer and Thisse (1994) examine the effect of commodity taxation in a setting in which firms sell vertically differentiated products and compete on prices. They show that an increase in taxes reduces the quality of the products that oligopolists will provide. Although we assume that the brands offered are fixed in the short run, our empirical results show that consumers purchase more subcompact and compact vehicles (lower quality products relative to large and midsize cars) when taxes increase. Thus the change in the tax rate leads to a demand shift. In the long run, the demand effect resulting from the tax change would likely encourage firms to reduce the number of ‘premium’ models they sell and offer more models in the subcompact and compact classes. In this sense, our results are in the spirit of Cremer and Thisse (1994); increases in taxes would lead to a shift in the distribution of sales to smaller and lower quality cars.

Empirical work in the literature has almost exclusively been conducted under

\(^2\)There are very few simulations that allow for both sides of the market to adjust. Most other simulations in this literature keep prices fixed.
the assumption that the relevant industry is competitive. Kenkel (1996) examines
the optimal rate of taxation on alcohol. In his setting, he assumes that there is a
single ‘alcoholic’ product that is priced at marginal cost. He acknowledges that the
“alcoholic beverage industry is probably better described as oligopolistic, but the
implications for tax incidence are unclear” (Kenkel, 1996, p. 300). To the best of
our knowledge, Barnett et al. (1995) is the only paper that empirically employs an
oligopoly model to examine the effect of taxation. Using a Cournot model with
homogeneous products to model competition among manufacturers in the cigarette
industry, they examine the incidence of cigarette taxes. Using the estimated
parameters, they simulate the effect of changes in the rate of cigarette taxation.
Given that there is not a great deal of vertical differentiation among the popular
cigarette brands, the homogeneous Cournot model is probably a reasonable model
to use in their setting.

In the automobile market in Israel, the price among the available models ranges
in 1995 from $14,833 for a Fiat Uno (a 1.1 litre engine with a standard
transmission and no air-conditioning) to $64,987 for a Volvo 960 (a 3.0 l engine,
air-conditioning, automatic transmission, airbags, ABS brakes, and other premium
features). Clearly, we cannot model competition in this industry by employing a
Cournot model with homogeneous products. We need a model that not only
captures total sales, but also the composition of the (differentiated) products sold
in the market.

2. The oligopolistic model

We model the automobile industry as an oligopolistic market with \( N \) multi-
product firms. Short run competition among the firms is through prices. Market
demand is determined by aggregating a discrete choice model of consumer
behavior.

2.1. Demand

Following Berry (1994), we use a random utility model, in which the utility of
product \( j \) to consumer \( i \), \( u_{ij} \), is:

\[
  u_{ij} = x_j \beta - \alpha p_j + \xi_j + e_{ij}
\]

(1)

where \( x_j \) is a vector of observed product characteristics (such as engine size) and \( p_j \)
is the price of automobile \( j \). \( \alpha \) and \( \beta \) are parameters to be estimated. The last two
terms of Eq. (1) are error terms: \( \xi_j \) is the average value of product \( j \)’s unobserved
characteristics (and is the same for all consumers) and \( e_{ij} \) represents the
distribution of consumer preferences around this mean. The term \( e_{ij} \) introduces
heterogeneity and its distribution determines the substitution patterns among
products. Under the assumption that the $\varepsilon_{ij}$ are identically and independently distributed across consumers and products with the extreme value (Weibull) distribution function, the probability of choosing product $j$ (the market share of product $j$) is given by:

$$s_j = \frac{e^{\delta_j}}{\left(\sum_k e^{\delta_k}\right)}$$

(2)

where

$$\delta_j = x_j \beta - \alpha \rho_j + \xi_j$$

(3)

is the mean utility level from product $j$. It is well known that this ‘logit’ distribution yields unreasonable substitution patterns among products; it is commonly used because it yields a closed form solution for the market share of each product.

In order to overcome the unreasonable substitution patterns, similar to other authors, we employ the ‘nested’ multinomial logit model. Goldberg (1995) and Verboven (1996) also employ variants of the nested logit model in their studies of the automobile industry.

Although, the nested logit model also has its limitations, it yields a more reasonable pattern of substitution among products than the logit model. Unlike the logit, the nested logit results in a higher degree of substitution among cars that belong to the same group than among cars from different groups.

As Berry (1994) notes, the nested logit is appropriate when the substitution effects among products depend primarily on predetermined classes of products. This assumption seems reasonable in the case of automobiles; indeed industry groups employ a standard classification system that puts each car in one of the following groups: subcompact, compact, midsize, large, and luxury/sport, according to its characteristics.

In the nested logit model, the products are grouped into $G+1$ sets, where the outside good, $j=0$, is assumed to be the only member of group 0. The difference between the logit and the nested logit is that there is an additional variable, denoted $\zeta$, which is common to all products in group $g$ and has a distribution that depends on $\sigma$, $0 \leq \sigma \leq 1$. In the case of the nested logit model, the utility of consumer $i$ from product $j$ is $u_{ij} = \delta_j + \zeta_j + (1-\sigma)\varepsilon_j$, where $\delta_j = x_j \beta - \alpha \rho_j + \xi_j$ is again the mean utility level and $\zeta + (1-\sigma)e$ also has an extreme value (Weibull) distribution. In this case, the probability of choosing product $j$ in group $g$ is:

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For more on general extreme value (GEV) models, see McFadden (1978).
\[ s_j = \frac{e^{\beta s_j(1-\sigma)}}{D_s^g(\sum D_s^{1-\sigma})} \] (4)

where \( D_s^g = \sum_{j \in G_s} e^{\beta s_j(1-\sigma)} \), \( G_s \) denotes the set of automobiles of type \( g \), and \( 0 \leq \sigma < 1 \) is an additional parameter to be estimated; it measures the degree of substitution among the products in the classes or groups. If \( \sigma = 0 \), the cross elasticities among products do not depend on the classification; in this case, the simple multinomial logit model of Eq. (1) is appropriate. When \( \sigma > 0 \), there is a higher degree of substitution among cars that belong to the same group than among cars from different groups. If \( \sigma \) approaches 1, the cross elasticity between any two cars that belong to different groups approaches 0.

We use the nested logit model to estimate the equilibrium in the Israeli automobile market. As shown in Berry (1994), Eq. (4) can be inverted to yield the following equation:

\[
\ln\left(\frac{s_j}{s_0}\right) = x_j \beta - \alpha p_j + \sigma \ln(\tilde{s}_{j|g}) + \xi_j
\] (5)

where \( \tilde{s}_{j|g} \) is the share of product \( j \) in group \( g \), and is the proportion of consumers that choose not to purchase a new car, that is, the proportion of consumers that choose the outside good. Following the literature, we assume that the size of the potential market (denoted by \( M \)) is known; \( s_0 \) is then the difference between the size of the potential market and the actual market. Since the price and the group share are endogenous, we can obtain consistent estimates of \( \alpha, \beta \) and \( \sigma \) from an instrumental variable regression.

2.2. Multiproduct oligopoly pricing

We assume that the marginal cost of producing each product is independent of the output levels and linear in a vector of cost characteristics. The assumption of constant marginal cost is typically employed in the relevant literature. Moreover, since the Israeli automobile market is small compared to the world market, the assumption of constant marginal cost is quite realistic in this case. The marginal cost of good \( j \) is:

\[
mc_j = \omega_j \gamma + \psi_j
\] (6)

where \( \omega_j \) is a vector of observable characteristics, \( \psi_j \) is an unobserved cost characteristic, and \( \gamma \) is a vector of unknown parameters to be estimated.

The operating profits of a multiproduct firm selling \( F \) different types of automobiles are:

\[
\pi_j = \sum_{k=1}^{F} \left( p_k / (1 + t) - mc_k \right) q_k
\] (7)
where \( p_k \) is the retail price of product \( k \), \( q_k \) is the corresponding quantity sold, \( t \) is the tax rate, and \( mc_k \) is the marginal cost of producing automobile \( k \).

We assume that firms compete through prices and that they only take into account the cross elasticities among their products within a group. It can be shown—with a lot of tedious algebra, see Verboven (1996)—that the first-order condition (pricing equation) for product \( j \) is:

\[
\frac{p_j}{1 + t} = \alpha_j \gamma + \frac{(1 - \sigma)}{\alpha(1 + t)(1 - \sigma)\sum_{k \in f_j} q_k/Q_g - (1 - \sigma)\sum_{k \in f_j} q_k/M} + \eta_j
\]

(8)

where \( f_j \) represents the set of products that firm \( f \) is selling in group \( g \), \( Q_f \) is the total number of sales in group \( g \), and \( M = \sum_{j \neq 0} q_j \). The last term on the right-hand side is endogenous, suggesting that instrumental variables are also needed in order to estimate the pricing equation.

3. Estimation

The model to be estimated consists of the demand and the pricing equations (Eqs. (5) and (8), respectively). We estimate this two equation system using the general method of moments (GMM). We chose to employ GMM estimation for the following reasons: (i) the unobserved demand characteristics, \( \xi_j \), and the unobserved cost characteristics, \( \upsilon_j \), might be correlated; (ii) \( \alpha \) and \( \sigma \) appear in both equations; (iii) the equations are not linear in \( \alpha \) and \( \sigma \). Additionally, other methods require structure on the correlation between the error terms of the demand and oligopoly pricing equation. GMM estimation does not require additional assumptions on the error term. Finally, GMM, which is an iterative procedure (like maximum likelihood), is preferable to 3SLS, which essentially involves a single iteration.

3.1. Instruments

We need to specify instruments for both the demand and the pricing equations. The endogenous terms for which we need instruments are product shares within a group (\( \bar{s}_{j|g} = q_j/Q_g \)), firm shares within a group, (\( \sum_{k \in f_j} q_k/Q_g \)), and prices. We use the characteristics of other cars and cost shifters as instruments.

The number of other products in a group and the sum of the characteristics of other products in a group are negatively correlated with within-group shares, and therefore can be used as instruments for this variable. Now consider firm shares within a group. This variable is positively correlated with the number of other products the firm sells in the group and with the sum of the characteristics of the
other cars it sells in the group. Further, firm shares within a group are negatively 
correlated with the number of products sold by competitors in the group, and with 
the sum of characteristics of products sold by competitors in the group. Finally, we 
consider instruments for price. The pricing equation suggests that an increase in 
the number of other automobiles that a firm sells within the group will increase the 
price. An important additional instrument for price is the change in the exchange 
rate between 1994 and 1995.⁴

3.2. Data

In the Israeli market the luxury/sport class is extremely small (due to the 
relatively high rate of taxation); hence we employ the classes: subcompact, 
compact, midsize and large. Approximately 113 000 automobiles were sold in both 
1994 and 1995. More than 170 different products were available in each year.⁵ 
Restricting the sample to brands that had more than 80 sales, left 213 brands: 101 

In Israel, all import licenses are exclusive, and the exclusive dealer sets prices. 
We used the Levi price book for price data, where prices are in New Israeli 
Shekels.⁶ The Levi price book, which is the most popular price book in Israel, 
includes the car features;⁷ hence for each price observation, we know what 
additional features were available⁸.

The retail price includes taxes of 144% on automobiles subject to custom duties, 
and 128% on automobiles not subject to custom duties. Total taxes are composed 
of the following three components: (i) a 95% luxury tax on private automobiles; 
(ii) a 17% value added tax; and (iii) a 7% customs tax; the taxes are cumulative. 
All private automobiles are subject to the luxury and value added tax. Automobiles 
that are produced and imported from the United States, Canada, and European 
Countries are exempt from custom duties because of free trade agreements. 
Automobiles from Japan and South Korea are not currently exempt from custom 
duties.

Our data includes the variable ENGINE, which is the engine size in l, and the 
dummy variables SUBCOMPACT, COMPACT, MIDSIZE and LARGE; these 
variables take on the value 1 if the automobile belongs to one of these classes. The 
dummy variables AIRCONDITION and AUTOMATIC take on the value 1 if the 
model has air conditioning or automatic transmission (respectively). AIRBRAKE

⁴Of course, all of the instruments are valid for all of the endogenous variables; we discuss the 
instruments in this manner so that we can provide the economic intuition behind the choice of 
instruments.

⁵Models with different engine size are different products.

⁶The average exchange rate in both 1994 and 1995 was 3.00 New Israeli Shekels = $1.00.

⁷In the Israeli market, many premium features (like dual airbags, automatic transmission, automatic 
braking system (ABS) etc.) are included as standard equipment or are not available.

⁸In the case in which options are available, we took the model with the fewest options.
takes on the value 2 if the model has both airbags and ABS brakes system, 1 if the model has only one of these features, and 0 if it has none of the features.

The dummy variable YEAR95 takes on the value 1 if the model was sold in 1995 and 0 if the model was sold in 1994. EXCHANGE takes on the value 0 if the model was sold in 1994 and equals the percentage change (from 1994 to 1995) in the exchange rate of the manufacturer’s country currency vs. the New Israeli Shekel, if the model was sold in 1995. The dummy variables JAPAN95, KOREA95, USA95, ITALY95 GERMANY95 and FRANCE95 take on the value 1 for 1995 automobiles that are produced in the relevant country. (These countries account for 85% of the automobiles sold in Israel in 1994–1995.) Descriptive statistics for these variables are shown in Tables A.1 and A.2 in Appendix A.

3.3. GMM estimation

In the estimation we include the variables ENGINE, AIRCONDITION, AUTOMATIC, AIRBRAKE both in the observable demand characteristics, \( x \), and the observable cost characteristics, \( \omega \). Additionally, the cost characteristics vector includes the variables EXCHANGE, YEAR95, JAPAN95, KOREA95, USA95, ITALY95 GERMANY95 and FRANCE95; the country dummy variables are included solely to examine the change in the marginal cost of production for each country from 1994 to 1995. (Country dummies for 1994 would have no meaning since EXCHANGE takes on the value 0 if the model is sold in 1994.) The estimated coefficients associated with the other variables are virtually unchanged if we exclude the country dummy variables. The demand characteristic vector includes a dummy variable for compact automobiles (denoted COMPACT); this variable is included because this class of vehicles is especially popular.

The instruments we employ are the sum of the engine sizes of the other products in the group, the sum of the engine sizes of the other products that a firm sells in the group, the number of other products in the group, the number of other products that a firm sells in the group, and the change in the exchange rate between 1994 and 1995. The results of the GMM estimation are shown Table 1.

The estimates of the marginal cost of air conditioning and automatic transmission are in line with the option prices that are listed separately in the Levi price book. (Recall that the prices are in New Israeli Shekels.)

The estimated elasticities of demand also seem quite plausible: the average (sales weighted) elasticity of demand is \(-4.80\); the breakdown is \(-3.53\) for the subcompact class, \(-4.70\) for the compact class, \(-5.62\) for the midsize class, and \(-9.15\) for the large class.

The model predicts that similar to the US, there is some market power in the Israeli automobile market: the average (sales weighted) price–cost margin is 10%. This is slightly lower than the price–cost margins obtained by BLP (1995) for the US automobile industry.

In an earlier version of the paper, we showed that with the exception of the
constant parameter in the demand equation, the parameter estimates are virtually unchanged when we significantly increase the size of the potential market from $M = 400 000$ to $M = 600 000$. As expected, the parameter estimate of the constant is significantly smaller when $M = 600 000$. A larger potential market ($M = 600 000$) means that more consumers chose the outside good than one of the available automobiles. This reduces the mean utility of all inside goods relative to the outside good.

### 3.4. Consumer surplus

Consumer surplus is calculated by noting that the equations in Eq. (4) are a system of probabilistic demand functions for individual $i$. This system exhibits all of the properties of deterministic demand functions. Therefore in the case of the nested logit model, consumer surplus per person (up to a constant) is given by: \(^9\)

$$W = \log\left(\frac{\sum_g D_g^{1-\sigma}}{\alpha}\right)$$

\(^9\)See McFadden (1978) for details.
Since we are interested in changes in consumer surplus associated with changes in taxation, the constant is not essential.

3.5. Simulation of the market

Given the parameters estimated in the previous section, in Sections 4 and 5 we simulate the equilibrium of the oligopolistic market equilibrium under different tax regimes. The importance of the simulation is that it allows firms to adjust prices, as well as allowing for changes in quantity demanded. Technically, each simulation involves solving 224 nonlinear equations, i.e. 112 demand equations [Eq. (5)] and 112 pricing equations [Eq. (8)] for each model without the error terms.\(^{10}\)

In the following section, we analyze the effects of changes in the ad valorem tax rate. In Section 5, we examine two alternative tax regimes.

4. The effect of changes in the ad valorem tax rate

We first examine the effect of the elimination of customs duties; these duties are only in effect for Japanese and Korean cars. This symmetry ensures that our results in Sections 4.1 and 4.2 below are not due to asymmetric taxation rates for automobiles that are identical except for country of origin.\(^{11}\)

We hence simulate the market under the assumption that all cars are subject to a 128% tax rate (the current tax rate without custom duties).\(^{12}\) Our market simulation indicates that eliminating custom duties on the Japanese and Korean cars increases sales by 2.3%, while government revenues from taxes decrease by 1.8%. The distribution among the four classes shifts from subcompact cars to compact cars, while the shares of the midsize and large classes stay almost the same. (This is because Japanese and Korean automobiles make up a very large portion of the compact class. See Table A.2 in Appendix A).

The elimination of the custom tax leads to an increase in the share of Japanese and Korean automobiles in each class while the share of European and American (E&US) decreases. Custom elimination implies a 6.6% reduction in the tax burden (for Japanese and Korean vehicles), where we define the tax burden to be the percentage change in prices such that the dealer price \(p/(1+t)\) is unchanged before and after the tax change. The simulation shows that Japanese and Korean

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\(^{10}\)This system was solved using software programs that we wrote; the software programs employed the GAUSS nonlinear simultaneous-equations subroutine.

\(^{11}\)Since Israel is now a signatory to the GATT, it suggests that the custom duties will be eliminated in the near future.

\(^{12}\)Because of the custom duties, the current tax on Japanese and Korean cars is 144%. 
firms would lower prices in the 4.1–5.8% range, following the elimination of custom duties. This implies that part of the tax reduction is passed onto consumers. The consumers’ share of the tax reduction is higher in the larger classes. Total consumer surplus increases by $46.9 million or 1.2%.  

Since the profits are all owned by foreign firms, the reduction in aggregate excess burden from the elimination of custom duties is the gain in consumer surplus less the decline in tax revenues or $21.9 million ($46.9 million less 25.0 million).

Even though it does not affect excess burden, the elimination of custom duties has a significant effect on the profits of the different foreign firms. The profits of American and European manufacturers fall by 4–10%, while the profits of the Japanese and Korean firms rise by anywhere from 20–33%. Honda has the largest increase in profits (33%); this is because their models are in the midsize and large classes. Thus despite the fact that the distribution of sales among classes does not change significantly, the distribution of sales (and profits) within classes changes significantly.

Table 2 summarizes the effects of eliminating customs duties:

4.1. A change in the ad valorem tax rate from 128% to 100%

Our basis for comparative statics in this section is the 128% uniform tax rate. Here we examine a tax reform that results in a decrease in the tax rate to 100%. Keeping the same tax regime enables us to examine the robustness of results to changes in parameters of the model.

Our simulations shows that the reduction in the uniform tax rate to 100% increases the market size by 10.4% relative to a 128% uniform tax; there is a rightward shift towards larger automobiles, since these cars are now relatively less expensive. See Table 3.

The economic reason for substitution towards larger cars when taxes fall from 128% to 100% is because the absolute tax decreases are larger in these classes.

Table 2
Changes following the elimination of custom duties

<table>
<thead>
<tr>
<th>Changes following the elimination of custom duties</th>
<th>Actual change</th>
<th>Percentage change</th>
</tr>
</thead>
<tbody>
<tr>
<td>Size of the market</td>
<td>+2408</td>
<td>+2.3</td>
</tr>
<tr>
<td>Total spending (millions $)</td>
<td>+4.6</td>
<td>+0.19</td>
</tr>
<tr>
<td>Tax revenues (millions $)</td>
<td>−25.0</td>
<td>−1.8</td>
</tr>
<tr>
<td>Consumer surplus (millions $)</td>
<td>+46.9</td>
<td>+1.2</td>
</tr>
<tr>
<td>Total firms’ profits (millions $)</td>
<td>+11.5</td>
<td>+5.1</td>
</tr>
</tbody>
</table>

13We use the equilibrium market size before the elimination of custom duties, although an argument could be made to use the predicted market size from the simulation following the elimination of the duties. We are consistent throughout the paper: we use the equilibrium market size before the relevant change in tax policy to compute the change in consumer surplus.
Table 3
Percentage change in prices and sales by class from decline in tax rate

<table>
<thead>
<tr>
<th>Fall in tax rate from 128% to 100%</th>
<th>Subcompact</th>
<th>Compact</th>
<th>Midsize</th>
<th>Large</th>
<th>Overall</th>
</tr>
</thead>
<tbody>
<tr>
<td>Percentage change in price</td>
<td>-8.6</td>
<td>-9.4</td>
<td>-9.9</td>
<td>-10.6</td>
<td>-9.4</td>
</tr>
<tr>
<td>Percentage change in sales</td>
<td>5.5</td>
<td>9.4</td>
<td>13.2</td>
<td>23.2</td>
<td>10.4</td>
</tr>
<tr>
<td>Overall elasticity of demand:</td>
<td>-1.1</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

While the size of the ‘shift’ is affected by the assumption that utility is linear in prices, qualitatively similar results would obtain even if utility were not linear in prices.

The reduction in the tax rate reduces the tax burden by 12.3%. (Recall that the tax burden is the percentage change in prices such that the dealer price \( p/(1+t) \) is unchanged before and after the tax change.) Table 3 shows that sales weighted prices fall by 9.4%; hence, consumers receive a significant share of the benefit from the reduction in the luxury tax. The percentage decline in prices is greater for larger cars.

Table 3 also shows the ‘total elasticity of automobile demand,’ which measures the relative demand for automobiles and the outside good. The total elasticity of demand is calculated by dividing the percentage change in overall sales (+10.4%) by the sales weighted percentage change in prices (−9.4%). Hence the total elasticity of automobile demand is approximately −1.1.

The size of the market increases by 11,169 (+10.4%) and consumer surplus increases by 238.4 million (+8.7%) following the decrease in the luxury tax. Government revenues from taxes fall by $120.0 million (−8.7%), while firms’ profits increase by $60.3 million (25.5%). Table 4 summarizes these results.

4.1.1. Robustness of results: changes in the size of the potential market

The estimates of tax incidence and overall elasticity of automobile demand are robust to changes in the size of the potential market. When the size of the potential market is 600,000 rather than 400,000 (i.e. 50% higher), the total elasticity of automobile demand is approximately −1.3. The slightly more elastic total

Table 4
Global changes from decline in tax rate from 128% to 100%

<table>
<thead>
<tr>
<th>Changes following decrease in the luxury tax from 128% to 100%</th>
<th>Actual changes</th>
<th>Percentage change</th>
</tr>
</thead>
<tbody>
<tr>
<td>Size of the market</td>
<td>+11,169</td>
<td>+10.4</td>
</tr>
<tr>
<td>Total spending (millions $)</td>
<td>+61.8</td>
<td>+2.5</td>
</tr>
<tr>
<td>Tax revenues (millions $)</td>
<td>−120.0</td>
<td>−8.7</td>
</tr>
<tr>
<td>Consumer surplus (millions $)</td>
<td>+238.4</td>
<td>+8.7</td>
</tr>
<tr>
<td>Total firms’ profits (millions $)</td>
<td>+60.3</td>
<td>+25.5</td>
</tr>
</tbody>
</table>
The elasticity of automobile demand is because the size of the (potential) market is larger when \( M = 600 \, 000 \). The only significant difference is in the absolute level of sales. At a tax rate of 100% and a potential market size of 400,000, predicted sales in the market are 119,034; at the same tax rate, when \( M = 600 \, 000 \), predicted sales are 114,262. This must be the case, because estimating the model with a larger potential market (but no other changes, i.e. tax rate stays the same) means that more consumers chose the outside good. As discussed above, this reduces the mean utility of all inside goods (automobiles) relative to the mean utility of the outside good and hence the simulation predicts that less automobiles will be purchased.

Thus, the size of the potential market is just a scaling factor and does not significantly affect tax incidence or the elasticity of demand between the inside and the outside goods. This is important, since, the potential market size is typically something that must be assumed. We continue the analysis with \( M = 400 \, 000 \).

4.1.2. Robustness of the results to changes in the degree of substitution between classes of automobiles (\( \sigma \))

When \( \sigma \) is relatively small, Eq. (8) shows that the mark-up over marginal cost is primarily a function of the firm’s share in the total market, while when \( \sigma \) is relatively large, the mark-up over marginal cost is primarily a function of the firm’s share within the relevant class of cars. When \( \sigma \) approaches 1, the price cost margin goes to 0, i.e. the market becomes very competitive; hence for larger values of \( \sigma \), we would expect firms to bear a larger portion of the tax burden.

We simulated the market for \( \sigma = 0.63 \) and \( \sigma = 0.77 \), i.e. one standard deviation below and one standard deviation above the estimated value (\( \sigma = 0.70 \)). On average, a reduction in the tax rate from 128% to 100%, leads to a price fall of 9.0% for the case when \( \sigma = 0.63 \) vs. 9.9% for the case when \( \sigma = 0.77 \). (Recall that in the case when \( \sigma = 0.70 \), the estimated value, prices fall on average by 9.4%.) The simulations with \( \sigma = 0.63 \) and \( \sigma = 0.77 \) show that the overall elasticity of demand for automobiles remains unchanged (−1.1) as \( \sigma \) changes.

4.1.3. Precision of results

In order to get some sense of the precision of the effect of the changes in taxation policy, we performed the following ‘bootstrap’ experiment for consumer surplus. It was assumed that the estimated coefficients from Table 1 were the true means and the estimated standard deviations were the true standard deviations of the unknown parameters. Further, we assumed that each of these parameters was normally distributed. Using Eq. (9), we then computed 5000 estimates of the change in consumer surplus. Fig. 1 shows that the estimated standard deviation of the aggregate gain in consumer surplus from the reduction in the tax rate from 128% to 100% is approximately $24.9 million.
Fig. 1. Aggregate gain in consumer surplus (in millions $) from reduction in the tax rate from 128% to 100%.

4.2. Benefits from estimating oligopolistic model with differentiated products

What are the benefits of using a setup with heterogeneous demand and oligopolistic market structure in estimating the effect of government taxation rather than the standard setup of a competitive market? To study this issue, we examine the change in government revenue and consumer surplus derived from a change in tax policy under the assumption that market is competitive and that the products are homogeneous. Specifically, we make the assumption that the percentage change in sales from the reduction in the tax rate from 128% to 100% is exactly as we measured using the oligopoly model. Then the change in government tax revenue \[\Delta pqt\] is simply \[pq \, dt + pt \, dq\]. The first term is given by total revenue in the market (at the 128% tax rate equilibrium) multiplied by the change in the tax rate, \((100 - 128)/228\); this yields $-301.4 million. The second term is total tax revenue at the 128% tax rate equilibrium multiplied by the change in sales (10.4%); this yields $143.3 million. Hence the ‘standard’ model estimates that government revenue falls by $158.1 million. Since the gain in consumer surplus is simply \[2pq \, dt\], the standard model estimates that reducing the tax rate from 128% to 100% results in an increase of $301.4 million in consumer surplus.

Table 4 shows that consumer surplus under the oligopoly model increases by $238.4 million when the tax rate is reduced from 128% to 100%. The difference between the two consumer surplus measures is $63 million, which is approximately the estimated value of the increase in firm profits ($60.3 million from Table 4).

\[\text{We are grateful to Roger Gordon for suggesting that we conduct the analysis in this section and for providing the framework for the analysis.}\]
when the tax rate decreases. Hence the oligopoly profit effect from the increase in demand (associated with the decrease in the tax rate) is fairly significant—approximately 25.3% of the increase in consumer surplus. This profit effect is not captured in the standard model.

Table 4 also shows that tax revenues fall by $120 million, rather than $158.1 million as predicted by the standard model. This difference is due to the increase in demand for and sales of large cars (see Table 3). Hence the standard model overestimates the reduction in government tax revenue and overestimates the gain in consumer surplus and these overestimates are fairly significant.

5. Alternative tax regimes: per car tax and tax based on engine size

In this section, we examine the effect of alternative taxation regimes. In particular, we consider: (i) a per unit tax; and (ii) a tax based on engine size. The effects of product differentiation and market power can be illustrated by comparing these taxation regimes to the (100%) uniform tax rate.

Note that both of these taxes are ‘additive’, rather than multiplicative. Hence, the first-order condition Eq. (7) becomes

\[ \pi_r = \sum_{k=1}^{r} (p_k - mc_k - T_k)q_k \]  

and the pricing Eq. (8) becomes

\[ p_j = T_j + \omega_j + \frac{(1 - \sigma)}{\alpha} \left\{ 1 - \sigma \left[ \sum_{k \in f_s} q_k / Q_k - (1 - \sigma) \sum_{k \in f_s} q_k / M \right] \right\} + u_j \]  

where \( T \) is the amount of the tax.

5.1. Per unit taxes

We first consider a per unit tax; such a tax is ‘regressive’ in the sense that the effective tax rate is higher on smaller vehicles. We examine a shift from a 100% uniform tax rate to a constant absolute tax of $10 500 per vehicle (the average tax per automobile in the case of a 100% uniform tax).\(^{15}\)

The effect of vertical product differentiation is illustrated by the change in the tax regime. This demand shift significantly affects the composition of the market: sales of subcompacts decline by 23.9% and sales of compacts decline by 13.4%.

\(^{15}\)When making regime changes, authorities are often concerned that tax revenue remains unchanged. It is straightforward to conduct this exercise. We chose to examine the per unit tax that leaves average tax per automobile equal to the previous regime because it illustrates the effects of market power and product differentiation clearly.
relative to sales under the 100% uniform tax rate. On the other hand, sales of midsize automobiles rise slightly (by 1.8%) and sales of large automobiles increase significantly (by 38.6%).

Of course, market power also affects sales. Firms take advantage of the increase in demand for larger cars. In the case of large cars, the change in regime leads to a per vehicle tax reduction of 36.2%. Prices on the other hand, only decline by 10.7%. The significant increase in demand for large cars provides firms with an incentive to raise prices. This offsets much of the percentage decline in taxes and, to some extent, mitigates the demand effect.

In the case of midsize cars, the final price to consumers actually increases by 6.3%, despite the 8.6% decline in taxes on midsize cars. Although prices increase, sales of midsize cars actually increase overall by 1.8%. This is because the demand effect (associated with the change in tax regime) more than offsets the higher prices.

For subcompact cars, the demand effect works in the opposite direction. Hence, overall prices rise by ‘only’ 34.8%, less than the percentage increase in taxes (37.6%) on these vehicles (see Table 5).

The change in tax regimes leads to a 7.8% decline in total sales from approximately 119,000 to slightly less than 110,000. Total consumer spending rises by 8.8%. This is again due to the demand effect associated with the change in tax regime. Consumers switch to larger cars, which have become relatively less expensive. Given the above discussion, it is not surprising that firm profits increase significantly (by 83.6%). This is intuitive because profits per automobile are greater for larger cars and many more large automobiles are sold (see Table 6).

5.2. Tax on engine size

Here, we examine a change in the tax regime from a 100% uniform tax to a tax on engine size, with the tax approximately equal to $6600 times the size of the

Table 5
Percentage change in prices and sales by class from changes in tax regimes (in both cases the comparison is made with a 100% uniform tax rate)

<table>
<thead>
<tr>
<th>Engine Size</th>
<th>Subcompact</th>
<th>Compact</th>
<th>Midsize</th>
<th>Large</th>
</tr>
</thead>
<tbody>
<tr>
<td>$6600 × engine size</td>
<td>21.2</td>
<td>16.3</td>
<td>11.5</td>
<td>4.8</td>
</tr>
<tr>
<td>Percentage change in price</td>
<td>-13.3</td>
<td>-12.8</td>
<td>-7.3</td>
<td>-0.8</td>
</tr>
<tr>
<td>Percentage change in sales</td>
<td>+10.1</td>
<td>+7.3</td>
<td>+1.6</td>
<td>-0.6</td>
</tr>
</tbody>
</table>

Per unit tax of $10,500 per vehicle

<table>
<thead>
<tr>
<th>Engine Size</th>
<th>Subcompact</th>
<th>Compact</th>
<th>Midsize</th>
<th>Large</th>
</tr>
</thead>
<tbody>
<tr>
<td>Percentage change in price</td>
<td>34.8</td>
<td>16.5</td>
<td>6.3</td>
<td>-10.7</td>
</tr>
<tr>
<td>Percentage change in sales</td>
<td>-23.9</td>
<td>-13.4</td>
<td>+1.8</td>
<td>+38.6</td>
</tr>
<tr>
<td>Percentage change in taxes</td>
<td>+37.6</td>
<td>+7.8</td>
<td>-8.6</td>
<td>-36.2</td>
</tr>
</tbody>
</table>
Table 6
Global changes associated with different taxation regimes (in both cases, the comparison is made with the base case, a 100% uniform tax rate)

<table>
<thead>
<tr>
<th>Changes relative to uniform tax of 100%</th>
<th>Per unit tax $10 500 per vehicle</th>
<th>Tax on engine size $6600×engine size (l)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Actual changes</td>
<td>Percentage change</td>
</tr>
<tr>
<td>Size of the market</td>
<td>−9227</td>
<td>−7.8</td>
</tr>
<tr>
<td>Total spending (millions $)</td>
<td>+222.6</td>
<td>+8.8</td>
</tr>
<tr>
<td>Tax revenues (millions $)</td>
<td>−98.1</td>
<td>−7.8</td>
</tr>
<tr>
<td>Consumer surplus (millions $)</td>
<td>−224.5</td>
<td>−6.8</td>
</tr>
<tr>
<td>Total firms’ profits (millions $)</td>
<td>+247.9</td>
<td>+83.6</td>
</tr>
</tbody>
</table>

This results in a tax of $10 500 (the average tax per vehicle in the case of a 100% uniform tax) on the average size automobile (1.6 l engine) in the market. This tax is still regressive relative to the uniform tax rate, but much less so than the very regressive constant tax. Hence the changes in prices and automobile sales by class are less dramatic than in the case of changing from a uniform tax rate to a constant tax per automobile.

Prices of all classes of automobiles rise, with the percentage increase in prices decreasing in the size of the car. Sales also decrease monotonically by class; sales of subcompacts fall by 13.3%, while sales of large cars fall just by 0.8%. Table 5 summarizes these changes, which again are due to a combination of product differentiation, market power, and the relative changes in taxes per automobile.

6. Concluding remarks

While the paper examines the effect of different tax regimes on the Israeli automobile market, the main contribution of this paper is the framework it adopts. As we have shown, in oligopolistic industries, taxation affects the profile of goods that are sold as well as relative prices in a way that depends on the elasticity of demand of all products and the degree of competition in the market.

Our approach could be used by regulators to predict the effect of different tax policies, not only on tax revenues and the number of cars that are sold in the market, but also on the distribution of the type of automobiles that are sold. Hence, different tax regimes can be evaluated and compared for their effect on the average engine size (which determines the consumption of fuel), the percentage of automobiles that offer safety features (such as airbags and ABS brakes), as well as the level of imports.

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16Differential tax rates based on engine size are employed in Europe. For example, both Belgium and Italy charge higher tax rates for automobiles with large engine sizes. (See Verboven, 1996.)
Appendix

Table A.1
Descriptive statistics

<table>
<thead>
<tr>
<th>Variable</th>
<th>Mean</th>
<th>Maximum</th>
<th>Minimum</th>
</tr>
</thead>
<tbody>
<tr>
<td>PRICE</td>
<td>68.481</td>
<td>194,962</td>
<td>29,999</td>
</tr>
<tr>
<td>QUANTITY</td>
<td>1044</td>
<td>11,447</td>
<td>80</td>
</tr>
<tr>
<td>ENGINE</td>
<td>1.63</td>
<td>3.8</td>
<td>1.00</td>
</tr>
<tr>
<td>AIRCONDITION</td>
<td>0.88</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>AUTOMATIC</td>
<td>0.21</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>AIRBAGS</td>
<td>0.12</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>ABS BRAKES</td>
<td>0.07</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>SMALL</td>
<td>0.25</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>COMPACT</td>
<td>0.52</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>MEDIUM</td>
<td>0.16</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>LARGE</td>
<td>0.07</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>JAPAN</td>
<td>0.34</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>KOREA</td>
<td>0.16</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>FRANCE</td>
<td>0.13</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>ITALY</td>
<td>0.12</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>USA</td>
<td>0.07</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>GERMANY</td>
<td>0.04</td>
<td>1.00</td>
<td>0.00</td>
</tr>
<tr>
<td>EXCHANGE</td>
<td>4.91</td>
<td>13</td>
<td>-3</td>
</tr>
<tr>
<td>YEAR95</td>
<td>0.5</td>
<td>1.00</td>
<td>0.00</td>
</tr>
</tbody>
</table>

Table A.2
Automobile sales by group (J&K, Japan and Korea)

<table>
<thead>
<tr>
<th></th>
<th>Subcompact</th>
<th>Compact</th>
<th>Midsize</th>
<th>Large</th>
<th>Total</th>
</tr>
</thead>
<tbody>
<tr>
<td>1994 Total sales</td>
<td>25,026</td>
<td>58,075</td>
<td>19,004</td>
<td>9,087</td>
<td>111,192</td>
</tr>
<tr>
<td>1994 Models</td>
<td>19</td>
<td>37</td>
<td>20</td>
<td>25</td>
<td>101</td>
</tr>
<tr>
<td>1994 J&amp;K sales</td>
<td>3,733</td>
<td>36,733</td>
<td>13,031</td>
<td>1,579</td>
<td>55,116</td>
</tr>
<tr>
<td>1994 J&amp;K models</td>
<td>5</td>
<td>11</td>
<td>10</td>
<td>6</td>
<td>32</td>
</tr>
<tr>
<td>1995 Total sales</td>
<td>31,155</td>
<td>57,129</td>
<td>15,381</td>
<td>7,614</td>
<td>111,279</td>
</tr>
<tr>
<td>1995 Models</td>
<td>33</td>
<td>33</td>
<td>27</td>
<td>19</td>
<td>112</td>
</tr>
<tr>
<td>1995 J&amp;K sales</td>
<td>2,241</td>
<td>47,763</td>
<td>47,444</td>
<td>891</td>
<td>55,639</td>
</tr>
<tr>
<td>1995 J&amp;K models</td>
<td>6</td>
<td>16</td>
<td>12</td>
<td>2</td>
<td>36</td>
</tr>
</tbody>
</table>

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References


